

On Modelling of Infant Mortality Rate in Nigeria with Exponentiated Cubic Transmuted Exponential Distribution

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Abstract: The idea of introducing extra parameters into the existing model in enhancing more flexibility is a giant stride in research. Transmutation map technique is one of the recent methods of introducing additional properties such as skewness, kurtosis and bimodality into the baseline distribution. In this article, a new exponentiated exponential distribution is developed using transmutation map. This model is referred to as exponentiated cubic transmuted exponential distribution (ECTED). The mathematical properties of the model which include survival function, hazard function, central and non- central moments, moment generating function and order statistics are established. The inherent parameters in the model are estimated using method of maximum likelihood estimation (MLE). The system of equations obtained is non-linear in parameters therefore non-linear optimization algorithms are implemented in R package. The distribution is used to model data on infant mortality rate in Nigeria. The performance of the subject model is compared with its baseline exponential distribution (ED), transmuted exponential distribution (TED), exponentiated transmuted exponential distribution (ETED) and cubic transmuted exponential distribution (CTED) using Akaike Information criterion (AIC), Corrected Akaike Information criterion (AICC) and Bayesian Information criterion (BIC). It is hope that this will serve as an alternative distribution in modelling complex real life data arising from various fields of human endeavors.

Keywords: Exponentiated, Cubic, Infant Mortality, Reliability Function, Hazard Rate Function, Parameter Estimation, Order Statistics, Transmutation

1. Introduction

Exponential distribution has a lot of applications in various fields of human endeavors. The areas which include reliability analysis, biological sciences, economics, actuarial science, medicines and finance. The model has enjoyed generalizations over the years which include exponentiated exponential distribution, transmuted exponential distribution, Kw-exponential distribution, exponentiated transmuted exponential distribution to mention a few. This is to introduce flexibility into the parent model to enhance better goodness of fit in data analysis.

It worth noting that many works have been done on the transmutation of baseline distributions. Shaw et al [1] suggested a quadratic transmutation map and this was applied to exponential, uniform and normal distributions. The same map has been applied by Aryal et al [2] on the transmuted extreme value distribution, Aryal et al [3] on transmuted Weibull distribution, Aryal et al [4] on transmuted log-logistic, Merovci et al [5] on generalized transmuted family

of distributions, Merovci et al [6] on transmuted Lindley-geometric Distribution, Merovci et al [7] on transmuted generalized Rayleigh distribution, Merovci et al [8] on transmuted Pareto distribution and Merovci et al [9] on transmuted Lindley distribution. AL-Kadim et al [10] on Cubic transmuted Weibull distribution, Granzotto et al [11] on cubic transmuted distributions and Rahman et al [12] generalized the work of Shaw et al [1]. Others include Adeyinka et al [13] on four parameters transmuted generalized distribution, Adeyinka et al [14] on transmuted half logistic distribution, Adeyinka et al [15] on transmuted type I generalized half logistic distribution, Adeyinka [16] transmuted logistic distribution, Adeyinka [17] on type I generalized transmuted logistic distribution.

The exponentiation of transmuted distribution has been considered by Ebraheim [18] on exponentiated transmuted weibull distribution and Al-Kadim et al [19] on exponentiated transmuted exponential distribution.

This work focuses on exponentiation of transmuted

distribution. A new transmuted exponential distribution will emerge having both shape and bimodality properties and it will be referred to as exponentiated cubic transmuted exponential distribution (ECTED).

2. Exponentiated Cubic Transmuted Exponential Distribution (ECTED)

If a random variable X has exponential distribution with cumulative distribution function (cdf) and probability density function (pdf) given by

$$G(x) = 1 - e^{-mx}, x > 0, m > 0 \quad (1)$$

and

$$F(x) = [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^a, x > 0, m > 0, a > 0 \quad (5)$$

and the corresponding pdf is obtained by differentiating (5) with respect to x and it is given by

$$f(x) = ame^{-mx}[(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}][1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{a-1}, \\ x > 0, m > 0, a > 0. \quad (6)$$

The probability density function in (6) is called exponentiated cubic transmuted exponential distribution. When $\lambda = 0$, $\beta = 0$ and $a = 1$ in (5) it reduces to the baseline exponential distribution (ED) in (1). When $\beta = 0$ and $a = 1$ in (5) it reduces to transmuted exponential distribution (TED). When $\beta = 0$ in (5) it reduces to

$$g(x) = me^{-mx}, x > 0, m > 0. \quad (2)$$

Using the cubic rank transmutation map in [12] given by

$$F(x) = (1 + \lambda)G(x) + (\beta - \lambda)G^2(x) - \beta G^3(x) \quad (3)$$

with transmutation parameters $|\lambda| \leq 1, |\beta| \leq 1$ and $-2 \leq \lambda + \beta \leq 1$.

exponentiation of (3) it gives

$$F(x) = [(1 + \lambda)G(x) + (\beta - \lambda)G^2(x) - \beta G^3(x)]^a, a > 0 \quad (4)$$

with shape parameter a .

The corresponding cdf and pdf of exponentiated cubic transmuted exponential distribution (ECTED) are given by

exponentiated transmuted exponential distribution (ETED) and when $a = 1$ in (5) it reduces to cubic transmuted exponential distribution (CTED).

The graphs of cdf and pdf of ECTED for some selected values of parameters a, m, β and λ are shown in Figure 1 and Figure 2.

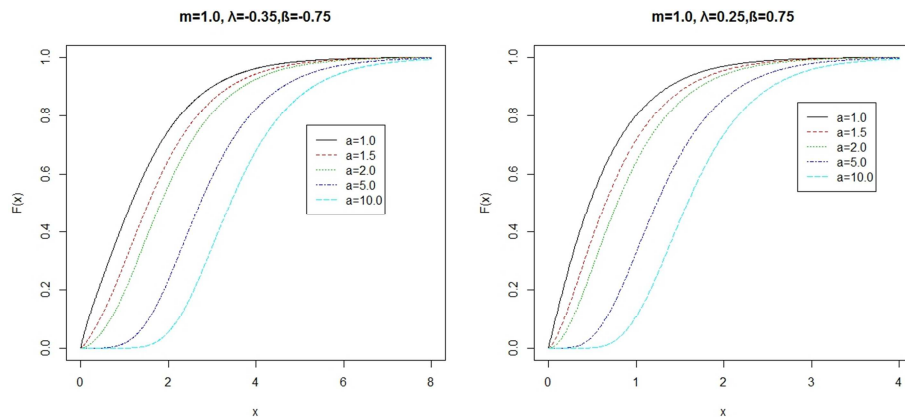


Figure 1. The graphs of cdf of ECTED for various values of parameters a, m, β and λ .

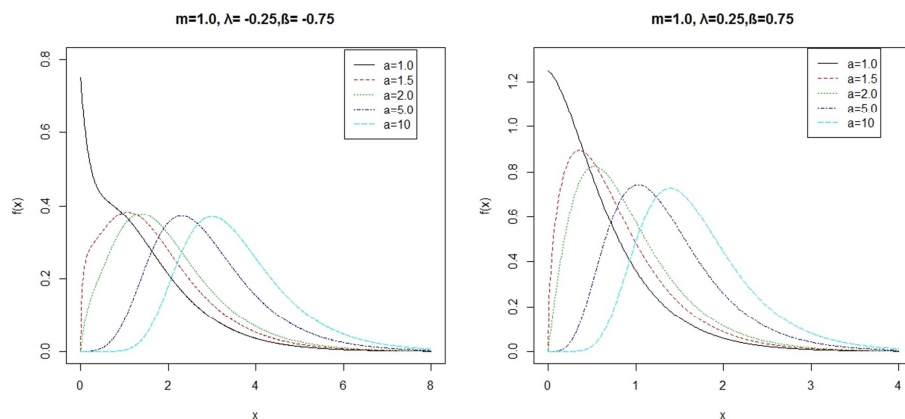


Figure 2. The graphs of pdf of ECTED for various values of parameters a, m, β and λ .

3. Central, Non-central Moments and Moment Generating Function

The p^{th} non-central moment of random variable X is given by

$$E[X^p] = am \int_0^\infty x^p e^{-mx} [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}] [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{a-1} dx. \quad (7)$$

Using binomial expansion

$$[1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{a-1} = \sum_{i=1}^{a-1} \sum_{j=0}^i \sum_{k=0}^j \binom{a-1}{i} \binom{i}{j} \binom{j}{k} (-1)^{j-k} \beta^k (\lambda + 2\beta)^{j-k} (\lambda + \beta - 1)^{i-j} \beta e^{-m(i+j+k)x}. \quad (8)$$

By substituting (8) into (7) it becomes

$$m \sum_{i=1}^{a-1} \sum_{j=0}^i \sum_{k=0}^j \binom{a-1}{i} \binom{i}{j} \binom{j}{k} (-1)^{j-k} \beta^k a (\lambda + 2\beta)^{j-k} (\lambda + \beta - 1)^{i-j} \int_0^\infty x^p e^{-m(i+j+k+1)x} [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}] dx. \quad (9)$$

The integral part of (9) gives

$$\frac{\Gamma(p+1)}{m^{p+1}} \left[\frac{1-\lambda-\beta}{(i+j+k+1)^{p+1}} + \frac{2(\lambda+2\beta)}{(i+j+k+2)^{p+1}} - \frac{3\beta}{(i+j+k+3)^{p+1}} \right] \quad (10)$$

By substituting (10) into (9)

$$E[X^p] = \frac{\Gamma(p+1)}{m^p} \sum_{i=1}^{a-1} \sum_{j=0}^i \sum_{k=0}^j \binom{a-1}{i} \binom{i}{j} \binom{j}{k} (-1)^{j-k} \beta^k a (\lambda + 2\beta)^{j-k} (\lambda + \beta - 1)^{i-j} \left[\frac{1-\lambda-\beta}{(i+j+k+1)^{p+1}} + \frac{2(\lambda+2\beta)}{(i+j+k+2)^{p+1}} - \frac{3\beta}{(i+j+k+3)^{p+1}} \right] \quad (11)$$

since $a - 1 > 0$.

The mean of ECTED is obtained by setting $p = 1$ in (11) to have

$$E[X] = \frac{1}{m} \sum_{i=1}^{a-1} \sum_{j=0}^i \sum_{k=0}^j \binom{a-1}{i} \binom{i}{j} \binom{j}{k} (-1)^{j-k} \beta^k a (\lambda + 2\beta)^{j-k} (\lambda + \beta - 1)^{i-j} \left[\frac{1-\lambda-\beta}{(i+j+k+1)^2} + \frac{2(\lambda+2\beta)}{(i+j+k+2)^2} - \frac{3\beta}{(i+j+k+3)^2} \right]. \quad (12)$$

The q^{th} central moment of ECTED is given by

$$E[x - \mu]^q = \sum_{p=0}^q \binom{q}{p} (-\mu)^{q-p} E[x^p] \quad (13)$$

By substituting (11) into (13) it gives

$$E[x - \mu]^q = \sum_{p=0}^q \sum_{i=1}^{a-1} \sum_{j=0}^i \sum_{k=0}^j \binom{a-1}{i} \binom{i}{j} \binom{j}{k} \binom{q}{p} (-1)^{j-k} \beta^k a (\lambda + 2\beta)^{j-k} (\lambda + \beta - 1)^{i-j} (-\mu)^{q-p} \frac{\Gamma(p+1)}{m^p} \left[\frac{1-\lambda-\beta}{(i+j+k+1)^{p+1}} + \frac{2(\lambda+2\beta)}{(i+j+k+2)^{p+1}} - \frac{3\beta}{(i+j+k+3)^{p+1}} \right]. \quad (14)$$

The variance of ECTED is obtained by setting $q = 2$ in (14) to have.

$$E[x - \mu]^2 = \sum_{p=0}^2 \sum_{i=1}^{a-1} \sum_{j=0}^i \sum_{k=0}^j \binom{a-1}{i} \binom{i}{j} \binom{j}{k} \binom{2}{p} (-1)^{j-k} \beta^k a (\lambda + 2\beta)^{j-k} (\lambda + \beta - 1)^{i-j} (-\mu)^{2-p} \frac{\Gamma(p+1)}{m^p} \left[\frac{1-\lambda-\beta}{(i+j+k+1)^{p+1}} + \frac{2(\lambda+2\beta)}{(i+j+k+2)^{p+1}} - \frac{3\beta}{(i+j+k+3)^{p+1}} \right]. \quad (15)$$

The moment generating function of ECTED is given by

$$M(t) = E[e^{tx}] = \sum_{p=0}^{\infty} \frac{t^p}{p!} E[x^p] \quad (16)$$

By substituting (11) into (16) it becomes

$$\sum_{p=0}^{\infty} \sum_{i=1}^{a-1} \sum_{j=0}^i \sum_{k=0}^j \binom{a-1}{i} \binom{i}{j} \binom{j}{k} (-1)^{j-k} \beta^k a (\lambda + 2\beta)^{j-k} (\lambda + \beta - 1)^{i-j} \frac{(m^{-1}t)^p \Gamma(p+1)}{p!} \left[\frac{1-\lambda-\beta}{(i+j+k+1)^{p+1}} + \frac{2(\lambda+2\beta)}{(i+j+k+2)^{p+1}} - \frac{3\beta}{(i+j+k+3)^{p+1}} \right]. \quad (17)$$

4. Reliability Analysis

The Survival function is given by

$$s(x) = 1 - F(x).$$

The survival function of ECTED is given by

$$s(x) = 1 - [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^a. \quad (18)$$

The graphs of survival function of ECTED for different values of parameters a , m , β and λ are shown in Figure 3.

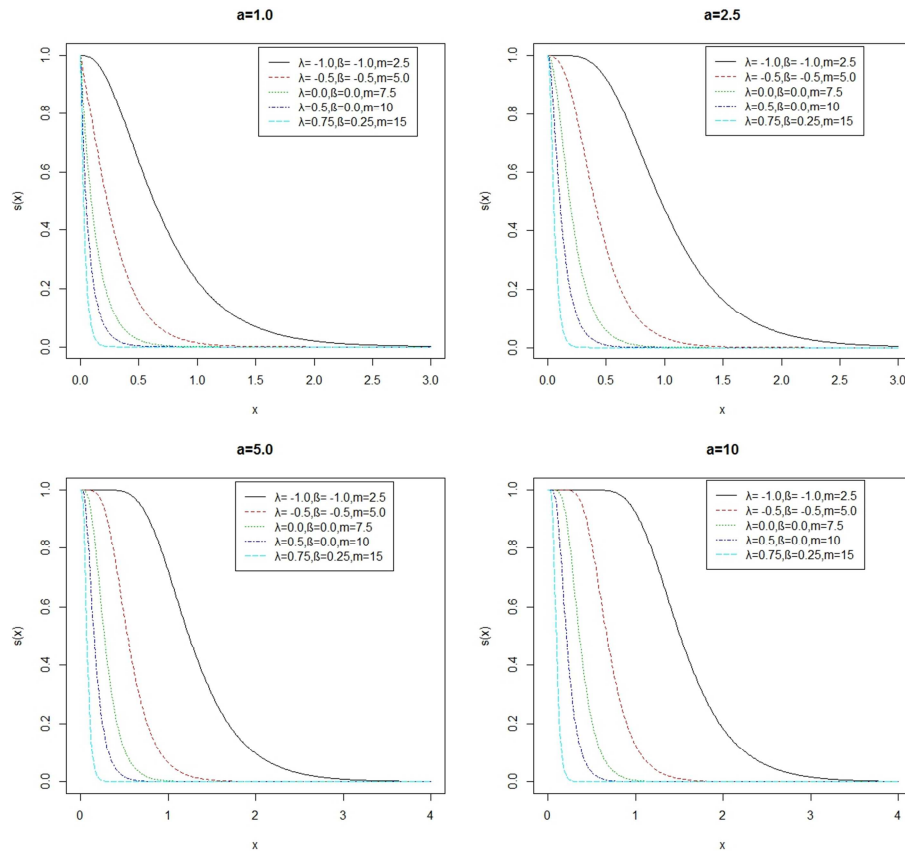


Figure 3. The graphs of survival function of ECTED for different parameters a , m , β and λ .

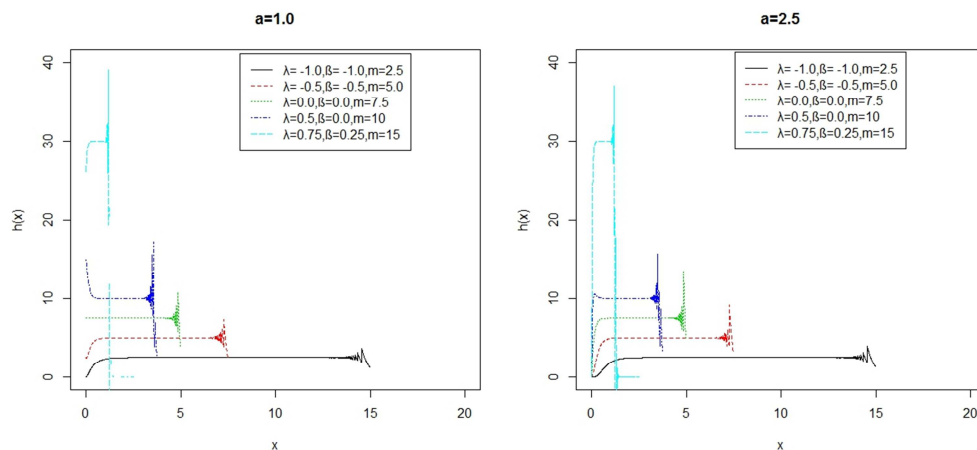
The hazard rate function is defined by

$$h(x) = \frac{f(x)}{s(x)}.$$

The hazard rate function of ECTED is given by

$$\frac{ame^{-mx}[(1-\lambda-\beta)+2(\lambda+2\beta)e^{-mx}-3\beta e^{-2mx}][1+(\lambda+\beta-1)e^{-mx}-(\lambda+2\beta)e^{-2mx}+\beta e^{-3mx}]^{a-1}}{1-[1+(\lambda+\beta-1)e^{-mx}-(\lambda+2\beta)e^{-2mx}+\beta e^{-3mx}]^a}. \quad (19)$$

The graphs of hazard rate function of ECTED for different values of parameters a , m , β and λ are shown in Figure 4.



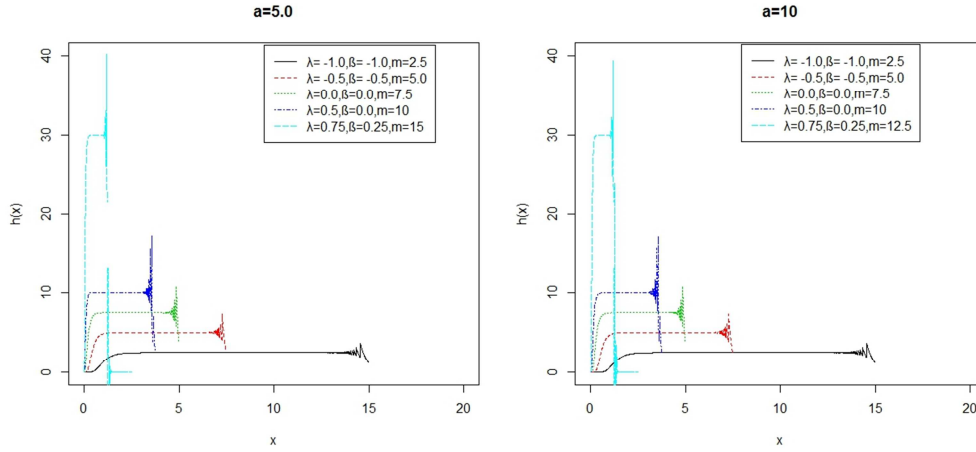


Figure 4. The graphs of hazard function of ECTED for different parameters a , m , β and $n\lambda$.

5. Parameter Estimation

Suppose X_1, X_2, \dots, X_n are independent random variables with sample size n from ECTED. Its likelihood function is given by

$$L = (am)^n e^{-m \sum_{i=1}^n x_i} \prod_{i=1}^n [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx_i} - 3\beta e^{-2mx_i}] \prod_{i=1}^n [1 + (\lambda + \beta - 1)e^{-mx_i} - (\lambda + 2\beta)e^{-2mx_i} + \beta e^{-3mx_i}]^{a-1}. \quad (20)$$

The log-likelihood function of (20) gives

$$\ln L = n \ln(am) - m \sum_{i=1}^n x_i + \sum_{i=1}^n \ln [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx_i} - 3\beta e^{-2mx_i}] + (a - 1) \sum_{i=1}^n \ln [1 + (\lambda + \beta - 1)e^{-mx_i} - (\lambda + 2\beta)e^{-2mx_i} + \beta e^{-3mx_i}]. \quad (21)$$

By taking partial derivative of (21) with respect to each parameter and equating the result to zero the following non-linear system of equations are obtained

$$\frac{\partial \ln L}{\partial a} = \frac{n}{a} + \sum_{i=1}^n \ln [1 + (\lambda + \beta - 1)e^{-mx_i} - (\lambda + 2\beta)e^{-2mx_i} + \beta e^{-3mx_i}] = 0 \quad (22)$$

$$\frac{\partial \ln L}{\partial m} = \frac{n}{m} - \sum_{i=1}^n x_i - 2 \sum_{i=1}^n \frac{(\lambda + 2\beta + 3\beta e^{-mx_i})x_i e^{-mx_i}}{[(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx_i} - 3\beta e^{-2mx_i}]} + (a - 1) \sum_{i=1}^n \frac{[(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx_i} - 3\beta e^{-2mx_i}]x_i e^{-mx_i}}{[1 + (\lambda + \beta - 1)e^{-mx_i} - (\lambda + 2\beta)e^{-2mx_i} + \beta e^{-3mx_i}]} = 0 \quad (23)$$

$$\frac{\partial \ln L}{\partial \lambda} = \sum_{i=1}^n \frac{2e^{-mx_i} - 1}{[(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx_i} - 3\beta e^{-2mx_i}]} + (a - 1) \sum_{i=1}^n \frac{e^{-mx_i}(1 - e^{-mx_i})}{[1 + (\lambda + \beta - 1)e^{-mx_i} - (\lambda + 2\beta)e^{-2mx_i} + \beta e^{-3mx_i}]} = 0 \quad (24)$$

$$\frac{\partial \ln L}{\partial \beta} = \sum_{i=1}^n \frac{4e^{-mx_i} - 3e^{-2mx_i} - 1}{[(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx_i} - 3\beta e^{-2mx_i}]} + (a - 1) \sum_{i=1}^n \frac{e^{-mx_i}(1 - 2e^{-mx_i} + e^{-2mx_i})}{[1 + (\lambda + \beta - 1)e^{-mx_i} - (\lambda + 2\beta)e^{-2mx_i} + \beta e^{-3mx_i}]} = 0 \quad (25)$$

The maximum likelihood estimator $\hat{\theta} = (\hat{a}, \hat{m}, \hat{\lambda}, \hat{\beta})'$ of parameters $\theta = (a, m, \lambda, \beta)'$ can be obtained by solving the system of equations given in (22), (23), (24) and (25).

A non-linear optimization algorithm such as quasi-Newton algorithm can be used to numerically maximize the log-likelihood function given in (21).

6. Order Statistics

Suppose $X_{(1)}, X_{(2)}, \dots, X_{(n)}$ are independent random variables which are arranged such that $X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$. This arrangement is called order statistics and it is denoted by $X_{(r)}$ with the cdf and pdf of a continuous population denoted by $F_X(x)$ and $f_X(x)$. David [20] gave the probability density function of $X_{(r)}$ as

$$f_{X_{(r)}}(x) = \frac{1}{B(r, n-r+1)} [F(x)]^{r-1} [1 - F(x)]^{n-r} f(x) \quad (26)$$

Such that $r = 1, 2, \dots, n$

The pdf of the r^{th} order statistics of ECTED which is obtained by substituting (5) and (6) into (26) is given by

$$f_{X_{(r)}} = \frac{amne^{-mx}}{B(r, n-r+1)} [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}] [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{ar-1} \{1 - [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^a\}^{n-r}. \quad (27)$$

Therefore the pdf of the minimum observation of $X_{(n)}$ is obtained by setting $r = 1$ in (27) and this gives

$$f_{X_{(1)}}(x) = amne^{-mx}[(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}][1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{a-1} \{1 - [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^a\}^{n-1}. \quad (28)$$

Using binomial expansion

$$\{1 - [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^a\}^{n-1} = \sum_{r=0}^{n-1} \binom{n-1}{r} (-1)^r [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{ar} \quad (29)$$

By substituting (29) into (28) it becomes

$$am \sum_{r=0}^{n-1} \binom{n-1}{r} (-1)^r ne^{-mx} [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}] [1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{ar+a-1}. \quad (30)$$

On expansion

$$[1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{ar+a-1} = \sum_{s=0}^{ar+a-1} \sum_{l=0}^s \sum_{w=0}^l \binom{ar+a-1}{s} \binom{s}{l} \binom{l}{w} (-1)^l \beta^w (\lambda + 2\beta)^{l-w} e^{-m(s+l+w)x} \quad (31)$$

On substitution of (31) into (30) it gives

$$m \sum_{r=0}^{n-1} \sum_{s=0}^{\infty} \sum_{l=0}^s \sum_{w=0}^l \binom{ar+a-1}{s} \binom{n-1}{r} \binom{s}{l} \binom{l}{w} (-1)^{r+l} \beta^w (\lambda + 2\beta)^{l-w} an e^{-m(s+l+w+1)x} [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}]. \quad (32)$$

The pdf of the maximum observation of $X_{(n)}$ is obtained by setting $r = n$ in (27) and it is given by

$$f_{X_{(n)}}(x) = amne^{-mx}[(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}][1 + (\lambda + \beta - 1)e^{-mx} - (\lambda + 2\beta)e^{-2mx} + \beta e^{-3mx}]^{an-1}. \quad (33)$$

Similarly, on expansion (33) gives

$$m \sum_{r=0}^{\infty} \sum_{s=0}^r \sum_{l=0}^s \binom{an-1}{r} \binom{r}{s} \binom{s}{l} (-1)^{s-l} \beta^l (\lambda + 2\beta)^{s-l} (\lambda + \beta - 1)^{r-s} an e^{-m(r+s+l+1)x} [(1 - \lambda - \beta) + 2(\lambda + 2\beta)e^{-mx} - 3\beta e^{-2mx}]. \quad (34)$$

7. Application

The data below represent the infant mortality rate per 1,000 live births in Nigeria between 1964 and 2018.

195.0, 190.7, 186.6, 182.4, 178.2, 173.7, 169.0, 164.2, 159.0, 154.0, 149.0, 144.1, 139.5, 135.4, 131.9, 128.8, 126.4, 124.7, 123.7, 123.3, 123.3, 123.7, 124.2, 124.7, 125.0, 125.2, 125.1, 124.9, 124.5, 124.0, 123.4, 122.3, 120.9, 118.9, 116.5, 113.7, 110.9, 107.8, 104.8, 101.6, 98.6, 95.6, 92.8, 90.2, 87.9, 85.9, 84.1, 82.7, 81.5, 80.5, 79.6, 78.7, 77.9, 76.9, 75.7.

Data source: www.data.unicef.org

The results from quasi Newton algorithm implemented in R package are shown in table 1. The table shows the performance of ECTED compared with ED, TED, ETED and

CTED using Akaike Information criterion (AIC), Corrected Akaike Information criterion (AICC) and Bayesian Information criterion (BIC).

$$AIC = 2k - 2LL, AICC = AIC + \frac{2k(k+1)}{n-k-1}$$

and

$$BIC = 2\log(n) - 2LL.$$

Where k is the number of parameters in the model, n is the sample size and LL is the maximized value of log likelihood function.

Table 1. The Results of the Analysis.

Model	Estimates	-LL	AIC	AICC	BIC
ED	$m=0.008$	309.42	620.84	620.92	626.85
TED	$m=1.469 \lambda=0.421$	302.51	609.02	609.25	613.04
ETED	$m=1.931 \lambda=0.496 a=1.274$	295.11	596.22	596.69	598.23
CTED	$m=2.451 \lambda=-0.211 \beta=-0.624$	291.63	589.26	589.73	591.27
ECTED	$m=3.015 a=1.492 \lambda=-0.256 \beta=-0.717$	275.87	559.74	560.54	559.75

It is observed from table 1 that ECTED performs better than other models.

8. Conclusion

A new exponentiated transmuted exponential model from cubic transmutation map which is called exponentiated cubic

transmuted exponential distribution (ECTED) is developed. The mathematical properties of the model are investigated. These properties include moments, moment generating function, order statistics, survival function and hazard rate function. The distribution is applied to model data on infant mortality rate in Nigeria. It is concluded that the subject model performs better than some well-known models in

literature. The results gotten from this work can be used in these field and other related disciplines. Further research can be made on the subject model especially in the area of applicability in other related fields.

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